

Supplemental Material

1 Derivation of the conditional distribution in Gibbs sampling of BBTM-II

To simplify the notations, we ignore the predefined parameters $\alpha, \beta, \boldsymbol{\eta}$ in the following equations. Moreover, when $e_i=0$, we let $z_i=0$. Thus, we have $P(z_i=0|e_i=0) = 1$ and $P(e_i=0|\mathbf{e}^{-i}, \mathbf{z}^{-i}, \mathbb{B}) = P(e_i=0, z_i=0|\mathbf{e}^{-i}, \mathbf{z}^{-i}, \mathbb{B})$.

Now we show the derivation of $P(e_i, z_i|\mathbf{e}^{-i}, \mathbf{z}^{-i}, \mathbb{B})$. Using the chain rule, $P(e_i, z_i|\mathbf{e}^{-i}, \mathbf{z}^{-i}, \mathbb{B})$ can be rewritten as:

$$P(e_i, z_i|\mathbf{e}^{-i}, \mathbf{z}^{-i}, \mathbb{B}) = \frac{P(\mathbf{e}, \mathbf{z}, \mathbb{B})}{P(\mathbf{e}^{-i}, \mathbf{z}^{-i}, \mathbb{B})} \propto \frac{P(\mathbb{B}|\mathbf{z})P(\mathbf{z}|\mathbf{e})P(e_i)}{P(\mathbb{B}^{-i}|\mathbf{z}^{-i})P(\mathbf{z}^{-i}|\mathbf{e}^{-i})}. \quad (1)$$

In Eq.(1), $P(\mathbb{B}|\mathbf{z})$ can be obtained by integrating out $\boldsymbol{\Phi} = \{\phi_0, \dots, \phi_K\}$:

$$\begin{aligned} P(\mathbb{B}|\mathbf{z}) &= \int P(\mathbb{B}|\mathbf{z}, \boldsymbol{\Phi})P(\boldsymbol{\Phi})d\boldsymbol{\Phi} \\ &= \int \left(\prod_{i=1}^{N_B} P(b_i|z_i, \phi_{z_i}) \right) P(\boldsymbol{\Phi})d\boldsymbol{\Phi} \\ &= \int \prod_{k=0}^K \left(\frac{\Gamma(W\beta)}{\Gamma(\beta)^W} \prod_{w=1}^W \phi_{k,w}^{n_{k,w}+\beta-1} d\phi_{\mathbf{k}} \right) \\ &= \left(\frac{\Gamma(W\beta)}{\Gamma(\beta)^W} \right)^K \prod_{k=0}^K \frac{\prod_{w=1}^W \Gamma(n_{k,w} + \beta)}{\Gamma(n_{k,\cdot} + W\beta)}, \end{aligned} \quad (2)$$

where $\Gamma(\cdot)$ is the standard Gamma function, $n_{k,w}$ is the number of times that word w assigned to topic k , and $n_{k,\cdot} = \sum_{w=1}^W n_{k,w}$.

$P(\mathbf{z}|\mathbf{e})$ can be obtained by:

$$\begin{aligned} P(\mathbf{z}|\mathbf{e}) &= \left(\prod_{i:e_i=0} P(z_i=0|e_i=0) \right) \left(\int \prod_{j:e_j=1} P(z_j|\boldsymbol{\theta})P(\boldsymbol{\theta})d\boldsymbol{\theta} \right) \\ &= \int \frac{\Gamma(K\alpha)}{\Gamma(\alpha)^K} \prod_{k=1}^K \theta_k^{n_k+\alpha-1} d\boldsymbol{\theta} \\ &= \frac{\Gamma(K\alpha)}{\Gamma(\alpha)^K} \frac{\prod_{k=1}^K \Gamma(n_k + \alpha)}{\Gamma(n_{\cdot} + K\alpha)}. \end{aligned} \quad (3)$$

where n_k ($k>0$) is the number of biterms assigned to bursty topic k , and $n_{\cdot} = \sum_{k=1}^K n_k$ is the total number of biterms assigned to bursty topics.

$P(\mathbb{B}^{-i}|\mathbf{z}^{-i})$ and $P(\mathbf{z}^{-i}|\mathbf{e}^{-i})$ can be worked out in the same way:

$$P(\mathbb{B}^{-i}|\mathbf{z}^{-i}) = \left(\frac{\Gamma(W\beta)}{\Gamma(\beta)^W} \right)^K \prod_{k=1}^K \frac{\prod_{w=1}^W \Gamma(n_{k,w}^{-i} + \beta)}{\Gamma(n_{k,\cdot}^{-i} + W\beta)}, \quad (4)$$

$$P(\mathbf{z}^{-i}|\mathbf{e}^{-i}) = \frac{\Gamma(K\alpha)}{\Gamma(\alpha)^K} \frac{\prod_{k=1}^K \Gamma(n_k^{-i} + \alpha)}{\Gamma(n_{\cdot}^{-i} - 1 + K\alpha)}. \quad (5)$$

Since the Gamma function satisfies $\Gamma(x+1) = x\Gamma(x)$, we obtain the final conditional distribution by replacing terms in Eq.(1) with those in Eqs.(2-5):

$$\begin{aligned}
P(e_i = 0 | \mathbf{e}^{-i}, \mathbf{z}^{-i}, \mathbb{B}) &\propto (1 - \eta_{b_i}) \cdot \frac{(n_{0,w_i,1}^{-i} + \beta)(n_{0,w_i,2}^{-i} + \beta)}{(n_{0,\cdot}^{-i} + W\beta)(n_{0,\cdot}^{-i} + 1 + W\beta)} \\
P(e_i = 1, z_i = k | \mathbf{e}^{-i}, \mathbf{z}^{-i}) &\propto \eta_{b_i} \cdot \frac{(n_k^{-i} + \alpha)}{(n_{\cdot}^{-i} + K\alpha)} \cdot \frac{(n_{k,w_i,1}^{-i} + \beta)(n_{k,w_i,2}^{-i} + \beta)}{(n_{k,\cdot}^{-i} + W\beta)(n_{k,\cdot}^{-i} + 1 + W\beta)}
\end{aligned}$$